

TAIL-DEPENDENCE PROPERTIES OF SOME NEW TYPES OF COPULA MODELS (PART II)

DIETMAR PFEIFER

Institut für Mathematik
Fakultät V
Carl-von-Ossietzky Universität Oldenburg
Germany
e-mail: dietmar.pfeifer@uni-oldenburg.de

Abstract

We continue the investigation of the tail-dependence behaviour of some new types of copula models, published recently in [5] and [6].

1. Introduction

For the sake of simplicity, we concentrate our investigations to the two-dimensional case. Let U_1, U_2 be standard random variables, i.e., they follow a uniform distribution over the interval $[0, 1]$ each. Let further T_1, T_2 be real continuous functions over $\mathbb{R}^2$ and $W_1 = T_1(U_1, U_2)$, $W_2 = T_2(U_1, U_2)$. If W_1, W_2 already follow a continuous uniform distribution over $[0, 1]$ each, then (W_1, W_2) is a representative of

Keywords and phrases: copula construction, tail dependence.

2020 Mathematics Subject Classification: 062H05, 062H20.

Received October 24, 2025; Accepted October 31, 2025

© 2025 Fundamental Research and Development International

a two-dimensional copula. Otherwise, $(V_1, V_2) := (F_1(W_1), F_2(W_2))$ is a representative of a two-dimensional copula if F_i denotes the continuous c.d.f. of W_i , $i = 1, 2$.

Of particular interest especially for financial markets or risk management is the tail dependence of copulas which was explicitly treated for dependence-of-unity copulas in [2], [3] and [4], and for the new approach in [6], which we shall continue here. The simplest definition of the coefficient λ_U of upper and λ_L of lower tail dependence is

$$\lambda_U = \lim_{t \uparrow 1} \frac{P(W_1 > F_1^{-1}(t), W_2 > F_2^{-1}(t))}{1 - t},$$

$$\lambda_L = \lim_{t \downarrow 0} \frac{P(W_1 \leq F_1^{-1}(t), W_2 \leq F_2^{-1}(t))}{t},$$

see, e.g., [1], Def. 7.36, p. 247.

In case that $F_1 = F_2$, i.e., W_1 and W_2 have same distribution, we also have

$$\lambda_U = \lim_{s \uparrow \infty} \frac{P(W_1 > s, W_2 > s)}{1 - F(s)}, \quad \lambda_L = \lim_{s \downarrow -\infty} \frac{P(W_1 \leq s, W_2 \leq s)}{F(s)}.$$

2. Particular Cases

Case 1. Here we consider the choice

$$W_1 = T_1(U_1, U_2) = U_1 + U_2, \quad W_2 = T_2(U_1, U_2) = \max(U_1, U_2).$$

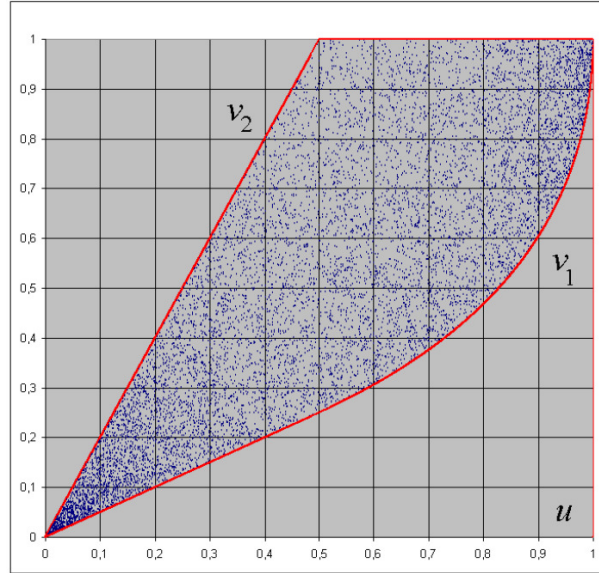
It is easy to see that the corresponding c.d.f.'s are given by

$$F_1(x) = \begin{cases} \frac{x^2}{2}, & 0 \leq x \leq 1, \\ 1 - 2\left(1 - \frac{x}{2}\right)^2, & 1 \leq x \leq 2 \end{cases}$$

and

$$F_2(x) = x^2, \quad 0 \leq x \leq 1.$$

The following graph shows 10,000 simulations of $(V_1, V_2) = (F_1(W_1), F_2(W_2))$.



The red lines (u, v) represent the (sharp) lower and upper envelopes of the copula, which are given by

$$v_1 = v_{lower} = \begin{cases} \frac{u}{2}, & \text{if } u \leq \frac{1}{2}, \\ \left(1 - \sqrt{\frac{1-u}{2}}\right)^2, & \text{otherwise} \end{cases}$$

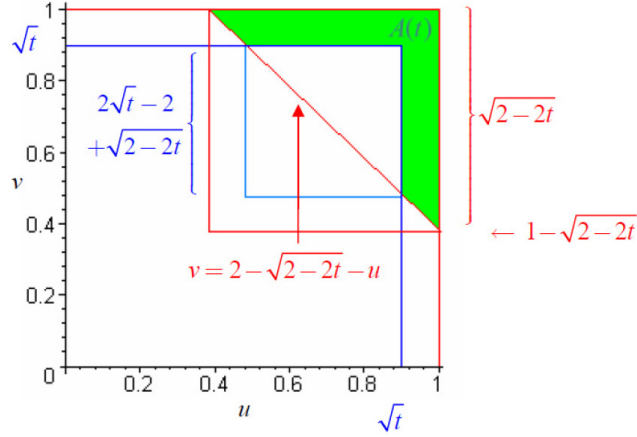
and

$$v_2 = v_{upper} = \begin{cases} 2u, & \text{if } u \leq \frac{1}{2}, \\ 1, & \text{if } u > \frac{1}{2}. \end{cases}$$

The lower bound is reached if V_1 and V_2 are close to each other, while the upper bound is reached if one of V_1 or V_2 is close to zero.

The subsequent graph explains our arguments for the calculation of the coefficient λ_U of upper tail dependence, which is given by $\lambda_U = 0$.

We start with some preliminary inequalities.



We have, for $t > \frac{1}{2}$, with μ denoting Lebesgue measure,

$$\begin{aligned}
 & P(W_1 > F_1^{-1}(t), W_2 > F_2^{-1}(t)) \\
 &= P(\max(U_1, U_2) > \sqrt{t}, U_2 > 2 - \sqrt{2-t} - U_1) \\
 &= \mu(A(t)) = \frac{1}{2} (\sqrt{2-2t})^2 - (2\sqrt{t} - 2 + \sqrt{2-2t})^2
 \end{aligned}$$

or, by a Taylor expansion around the point $t = 1$,

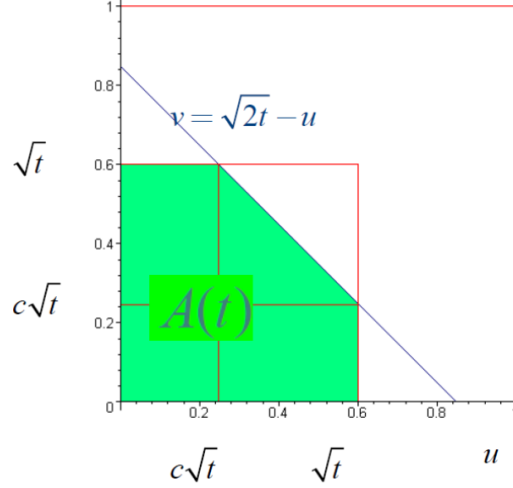
$$\mu(A(t)) = \sqrt{2}(1-t)^{3/2} + \mathcal{O}((1-t)^2), \text{ hence } \lambda_U = \lim_{t \rightarrow 1} \frac{\mu(A(t))}{1-t} = 0,$$

as stated.

The subsequent graph explains our arguments for the calculation of

the coefficient λ_L of lower tail dependence, which is given by $\lambda_L = 2(\sqrt{2} - 1) = 0.828427\dots$.

We start again with some preliminary inequalities.



We have, for $t > \frac{1}{2}$, with $c = \sqrt{2} - 1$,

$$\begin{aligned} & P(\max(U_1, U_2) \leq F_2^{-1}(t), U_1 + U_2 \leq F_1^{-1}(t)) \\ &= P(\max(U_1, U_2) \leq \sqrt{t}, U_2 \leq \sqrt{2t} - U_1) \\ &= \mu(A(t)) = t - \frac{\{(1-c)\sqrt{t}\}^2}{2} = t \left(1 - \frac{(1-c)^2}{2} \right) = 2ct \end{aligned}$$

and hence $\lambda_L = \lim_{t \downarrow 0} \frac{\mu(A(t))}{t} = 2c = 2(\sqrt{2} - 1) = 0.828427\dots$, as stated.

Case 2. Here we consider the choice

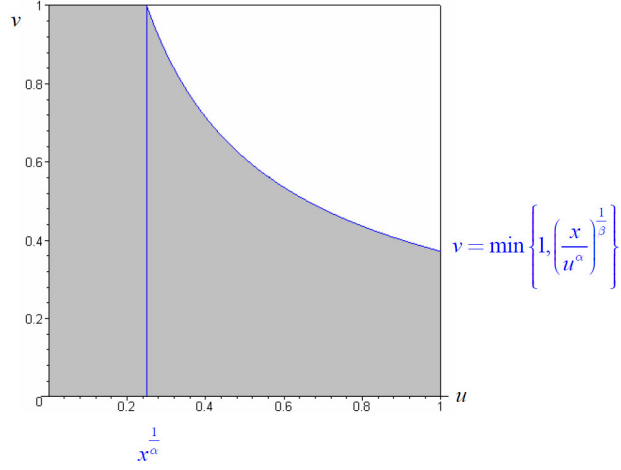
$$W_1 = T_1(U_1, U_2) = U_1^\alpha U_2^\beta, \quad W_2 = T_2(U_1, U_2) = U_1^\beta U_2^\alpha \quad \text{with real } \alpha, \beta > 0.$$

For simplicity, let $U := U_1$, $V := U_2$.

The common c.d.f. of U and V is given by

$$F(x) = \frac{\alpha}{\alpha - \beta} x^{\frac{1}{\alpha}} - \frac{\beta}{\alpha - \beta} x^{\frac{1}{\beta}}, \quad 0 < x < 1$$

which can be seen as follows.

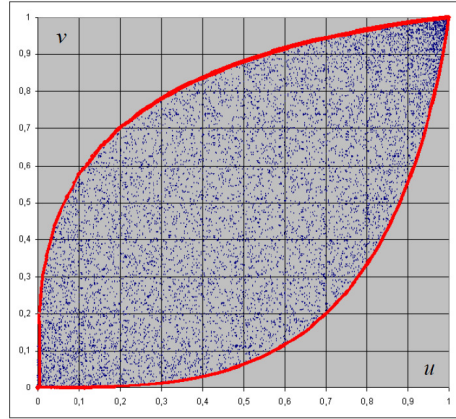


For $0 < x < 1$, there holds

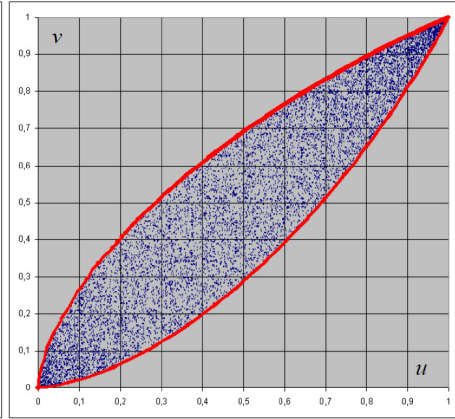
$$\begin{aligned} F(x) &= P(U^\alpha V^\beta \leq x) = P\left(V \leq \left(\frac{x}{U^\alpha}\right)^{\frac{1}{\beta}}\right) \\ &= \int_0^1 P\left(V \leq \left(\frac{x}{u^\alpha}\right)^{\frac{1}{\beta}}\right) du = \int_0^1 \min\left\{1, \left(\frac{x}{u^\alpha}\right)^{\frac{1}{\beta}}\right\} du \\ &= x^{\frac{1}{\alpha}} + x^{\frac{1}{\beta}} \int_{\frac{1}{x^\alpha}}^1 \frac{1}{u^{\frac{\alpha}{\beta}}} du = x^{\frac{1}{\alpha}} + \frac{\beta}{\beta - \alpha} x^{\frac{1}{\beta}} \left[1 - x^{\left(\frac{1}{\alpha} - \frac{1}{\beta}\right)}\right] \\ &= \frac{\alpha}{\alpha - \beta} x^{\frac{1}{\alpha}} - \frac{\beta}{\alpha - \beta} x^{\frac{1}{\beta}} = P(U^\beta V^\alpha \leq x) \end{aligned}$$

by symmetry reasons.

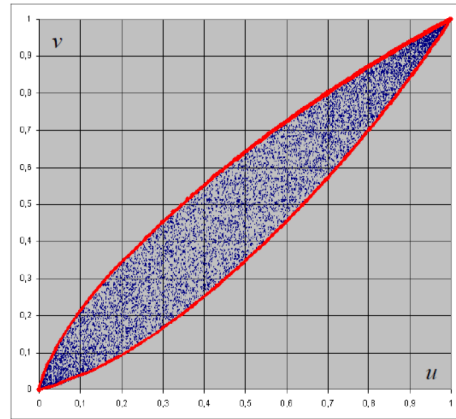
The following graphs show 10,000 simulations each of the copula given by $(F(W_1), F(W_2))$, for different values of α and β .



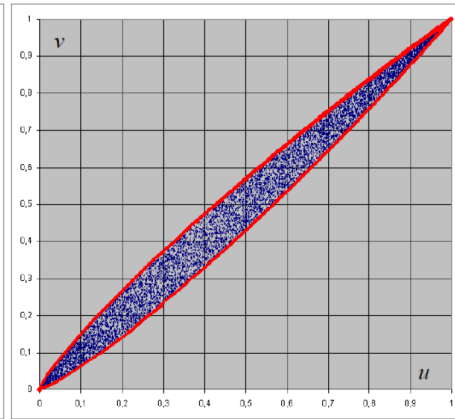
$\alpha = 3, \beta = 1$



$\alpha = 3, \beta = 2$



$\alpha = 4, \beta = 3$



$\alpha = 4, \beta = 3.5$

The red lines (u, v) represent the (sharp) lower and upper envelopes of the copula, which are given by

$$v_{lower} = F\left((F^{-1}(u))^{\frac{\beta}{\alpha}}\right) \quad \text{and} \quad v_{upper} = F\left((F^{-1}(u))^{\frac{\alpha}{\beta}}\right), \quad 0 < u < 1.$$

Alternatively, the lower envelope can be described by the points

$$\left[F(u), F\left(u^{\frac{\beta}{\alpha}}\right) \right] \text{ and the upper envelope by the points } \left[F(u), F\left(u^{\frac{\alpha}{\beta}}\right) \right],$$

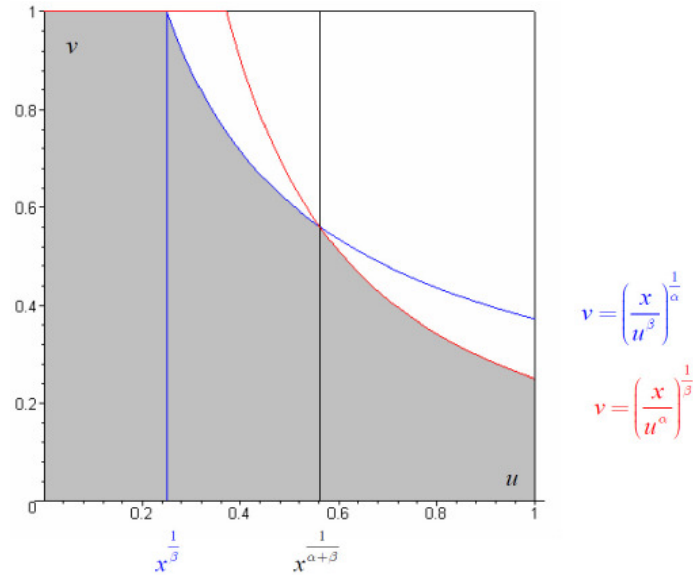
$$0 < u < 1.$$

Note that for $\alpha \gg \beta$, the copula tends to the independence copula, and for $\alpha \approx \beta$, we obtain the upper Fréchet bound. Note also that the copula is symmetric in α, β .

The subsequent graph explains our arguments for the calculation of the coefficient λ_L of lower tail dependence, which is given by $\lambda_L = 0$.

First notice that for $0 < x < 1$, the intersection point of $\left(\frac{x}{u^\beta}\right)^{\frac{1}{\beta}}$ and

$\left(\frac{x}{u^\alpha}\right)^{\frac{1}{\alpha}}$ is given by $u_x = x^{\frac{1}{\alpha+\beta}}$ with value u_x .



Next we have

$$\begin{aligned}
& P(U^\alpha V^\beta \leq x, U^\beta V^\alpha \leq x) \\
&= x^{\frac{1}{\beta}} + \frac{1}{x^{\frac{1}{\alpha+\beta}}} \int_{\frac{1}{x^{\frac{1}{\beta}}}}^{\frac{1}{x^{\frac{1}{\alpha+\beta}}}} \left(\frac{x}{u^\beta}\right)^{\frac{1}{\alpha}} du + \int_{\frac{1}{x^{\frac{1}{\alpha+\beta}}}}^{\frac{1}{x^{\frac{1}{\beta}}}} \left(\frac{x}{u^\alpha}\right)^{\frac{1}{\beta}} du = x^{\frac{2}{\alpha+\beta}}
\end{aligned}$$

with

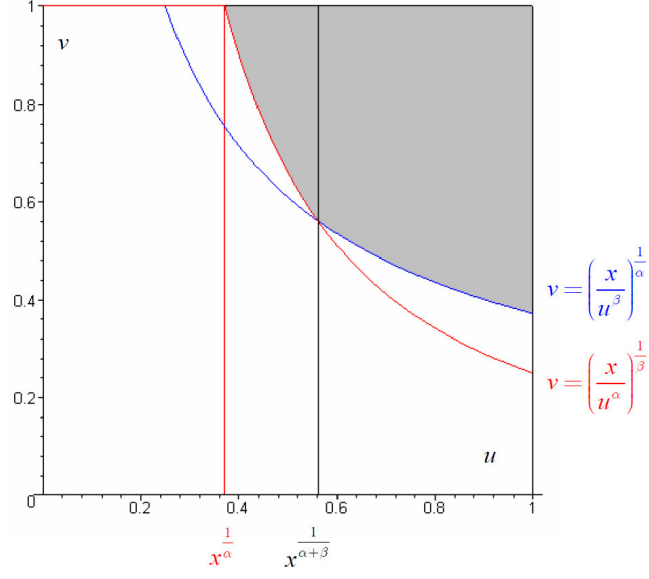
$$\frac{F(x)}{P(U^\alpha V^\beta \leq x, U^\beta V^\alpha \leq x)} = \frac{\alpha}{\alpha - \beta} x^{\frac{\beta-\alpha}{\alpha \cdot (\alpha+\beta)}} - \frac{\beta}{\alpha - \beta} x^{\frac{\alpha-\beta}{\alpha \cdot (\alpha+\beta)}}$$

and hence

$$\begin{aligned}
\lim_{x \rightarrow 0} \frac{F(x)}{P(U^\alpha V^\beta \leq x, U^\beta V^\alpha \leq x)} &= \infty, \text{ i.e.,} \\
\lim_{x \rightarrow 0} \frac{P(U^\alpha V^\beta \leq x, U^\beta V^\alpha \leq x)}{F(x)} &= 0 = \lambda_L.
\end{aligned}$$

The subsequent graph explains our arguments for the calculation of the coefficient λ_U of upper tail dependence, which is given by

$$\lambda_U = \frac{2\beta}{\alpha + \beta} > 0 \text{ if } \alpha > \beta.$$



If $\alpha > \beta$,

$$\begin{aligned}
 P(U^\alpha V^\beta > x, U^\beta V^\alpha > x) &= 1 - x^{\frac{1}{\alpha}} - \int_{\frac{1}{x^\alpha}}^{\frac{1}{x^{\alpha+\beta}}} \left(\frac{x}{u^\alpha}\right)^{\frac{1}{\beta}} du - \int_{\frac{1}{x^{\alpha+\beta}}}^1 \left(\frac{x}{u^\beta}\right)^{\frac{1}{\alpha}} du \\
 &= 1 - \frac{2\alpha}{\alpha - \beta} x^{\frac{1}{\alpha}} + \frac{\beta}{\alpha - \beta} x^{\frac{2}{\alpha+\beta}}
 \end{aligned}$$

with, by a Taylor expansion around $x = 1$,

$$\frac{P(U^\alpha V^\beta > x, U^\beta V^\alpha > x)}{1 - F(x)} = \frac{2\beta}{\alpha + \beta} - \frac{2(\alpha - \beta)}{3(\alpha + \beta)^2} (x - 1) + \mathcal{O}((x - 1)^2),$$

hence $\lambda_U = \frac{2\beta}{\alpha + \beta} > 0$.

References

- [1] A. J. McNeill, R. Frey and P. Embrechts, Quantitative Risk Management: Concepts, Techniques and Tools, rev. ed., Princeton University Press, New Jersey, 2015.
- [2] D. Pfeifer, H. A. Tsatedem, A. Mändle and C. Girschig, New copulas based on general partitions-of-unity and their applications to risk management, *Depend. Model.* 4 (2016), 123-140.
- [3] D. Pfeifer, A. Mändle and O. Ragulina, New copulas based on general partitions-of-unity and their applications to risk management (part II), *Depend. Model.* 5 (2017), 246-255.
- [4] D. Pfeifer, A. Mändle O. Ragulina and C. Girschig, New copulas based on general partitions-of-unity (part III) - the continuous case, *Depend. Model.* 7 (2019), 181-201.
- [5] D. Pfeifer, Reflections on a canonical construction principle for multivariate copula models, *Fundamental J. Math. Math. Sci.* 19(1) (2025), 97-107.
- [6] D. Pfeifer, Tail-dependence properties of some new types of copula models, *Fundamental J. Math. Math. Sci.* 19(1) (2025), 109-121.